

Finite Waiting Space Bulk Service System

V. P. SINGH

IBM Components Division, East Fishkill Facility, Hopewell Junction, N.Y. 12533, U.S.A.

(Received May 26, 1970)

SUMMARY

This paper discusses the ergodic queue length distribution of a bulk service system with finite waiting space by the method of the imbedded Markov chain. The system under consideration is a queuing system with Poisson arrivals, general service times, single server and where service is performed on batches of random size.

1. Introduction

In this paper we analyze a queuing system with Poisson arrivals, general service times, single server with variable batch capacity and finite waiting space. Such a system is denoted by $M/G^{(Y)}/1/(N+1)$ and is usually called a bulk service system. The stationary queue size distribution for this bulk service system is obtained by the method of the imbedded Markov chain. The method is based upon the fact that if the input is of the Poisson type, the length of the queue at epochs just when a batch departs constitutes a Markov chain which is "imbedded" in the continuous time parameter process. For an epoch at which there are no customers remaining to be served, the service is discontinued and the next observation on the process occurs at the end of the service period for the first subsequent arrival. The "imbedded" Markov chain is completely characterized once the transition probability matrix has been determined.

Special cases of the bulk service system considered in this paper have been studied by several authors such as Finch [3], Jain [4], Bhat [1,2], Takamatsu [9], Jaiswal [5], Rao [7], Keilson [6], Singh [8].

2. $M/G^{(Y)}/1/(N+1)$ Queuing System:

This bulk service system can be described as follows:

(i) Customers arrive one at a time in a Poisson process with parameter λ . The probability that j customers arrive in a time interval $(0,t)$ is therefore

$$k_j(t) = e^{-\lambda t} \frac{(\lambda t)^j}{j!}$$

(ii) The customers are served in batches of variable capacity. We assume the service capacity to be s , i.e., not more than s customers can be served at any time. Let t_1, t_2, \dots be the instants of departure of the successive batches, and denote by v_n the service time of the batch departing at t_n . We assume that $\{v_n\}$ is a sequence of identically distributed independent random variables with a common distribution function $G(t)$, $(0 \leq t < \infty)$. We further assume that v_n are independent of the arrival process.

(iii) Let $s - Y_n$ be the capacity for service ending at t_{n+1} ($n=0, 1, 2, \dots$). We assume that the random variables Y_n are identically distributed and mutually independent, and also independent of the arrival process. Let

$$\Pr \{ Y_n = r \} = b_r, \quad 0 \leq r \leq s, \\ = 0 \quad r > s$$

This may also happen if Y_n customers of the service ending at t_n need to go for another service, the total capacity of service of any batch being s . Then for the service starting after t_n , the server takes $\min(s - Y_n, \text{whole queue length})$. Let

$$B_j = \sum_{r=0}^j b_r = \Pr\{\text{customers already present with the server at a service epoch} \leq j\}$$

and

$$B_s(x) = \sum_{r=0}^s b_r x^r .$$

Note that

$$B_s(1) = B_s = \sum_{r=0}^s b_r = 1 \text{ and } B_0 = b_0 .$$

(iv) The waiting room has a fixed capacity of $(N + 1)$ customers (including those in service). An arrival finding the waiting room full balks and an arrival after joining the system does not renege. The server is never idle in the presence of customers.

3. Analysis of the $M/G^{(Y)}/1/(N + 1)$ System:

We say that the system is in state E_j when there are j customers in the system (including the batch just moving to be served). We define the transition probabilities:

$$\gamma_{ij} = \Pr\{\text{next state is } E_j \mid \text{previous state was } E_i\}$$

and the equilibrium probabilities

$$P_j = \Pr\{\text{the system is in state } E_j\}, \quad j = 0, 1, 2, \dots, N.$$

We next introduce the generating function

$$P(x) = \sum_{j=0}^{N-1} p_j x^j \tag{1}$$

Now let x_n be the number of customers arrived during a service period ending at t_n ; the distribution of x_n is given by

$$\Pr\{X_n = j\} = k_j = \int_0^\infty e^{-\lambda t} \frac{(\lambda t)^j}{j!} dG(t) \tag{2}$$

We next define

$$K(X) = \sum_{j=0}^{N-1} k_j x^j \tag{3}$$

The transition probability matrix for this system is given in Table 1.

From the transition probability matrix, we see that the equations determining the equilibrium probabilities are:

$$\begin{aligned} p_j = & k_j p_0 + (k_j B_{s-1} + k_{j-1} b_s) p_1 + (k_j B_{s-2} + k_{j-1} b_{s-1} + k_{j-2} b_s) p_2 + \dots \\ & + (k_j B_1 + k_{j-1} b_2 + \dots + k_{j-s+1} b_s) p_{s-1} \\ & + (k_j b_0 + k_{j-1} b_1 + \dots + k_{j-s} b_s) p_s + \dots \\ & + (k_{j-N+s-1} b_0 + k_{j-N+s} b_1 + \dots + k_{j-N+1} b_s) p_{N-1} \\ & + (k_{j-N+s} b_0 + k_{j-N+s-1} b_1 + \dots + k_{j-N} b_s) p_N \end{aligned}$$

for $j = 0, 1, 2, \dots, N - 1$ (4)

and

$$\begin{aligned} p_N = & l_N p_0 + (l_N B_{s-1} + l_{N-1} b_s) p_1 + \dots + (l_{s+1} b_0 + l_s b_1 + \dots + l_1 b_s) p_{N-1} \\ & + (l_s b_0 + \dots + l_0 b_s) p_N , \end{aligned}$$

where

TABLE 1
Transition probability matrix for $M/G^X/1/(N+1)$ system

$i \backslash j$	0	1	2	j	$N-1$	N
0	k_0	k_1	k_2	k_j	k_{N-1}	l_N
1	$k_0 B_{s-1}$	$k_1 B_{s-1} + k_0 b_s$	$k_2 B_{s-1} + k_1 b_s$	$\dots k_j B_{s-1} + k_{j-1} b_s$	$\dots k_{N-1} B_{s-1} + k_{N-2} b_s$	$l_N B_{s-1} + l_{N-1} b_s$
2	$k_0 B_{s-2}$	$k_1 B_{s-2} + k_0 b_{s-1}$	$k_2 B_{s-2} + k_1 b_{s-1} + k_0 b_s$	$\dots k_j B_{s-2} + k_{j-1} b_{s-1} + k_{j-2} b_s$	$\dots k_{N-1} B_{s-2} + k_{N-2} b_{s-1} + k_{N-3} b_s$	$l_N B_{s-2} + l_{N-1} b_{s-1}$
\vdots	\vdots	\vdots	\vdots	\dots	\vdots	\vdots
$s-1$	$k_0 B_1$	\vdots	\vdots	$\dots k_j B_1 + k_{j-1} b_2 + \dots + k_{j-s+1} b_s$	\vdots	\vdots
s	$k_0 b_0$	$k_1 b_0 + k_0 b_1$	$k_2 b_0 + k_1 b_1 + k_0 b_2$	$\dots k_j b_0 + k_{j-1} b_1 + \dots + k_{j-2} b_s$	$\dots k_{n-1} b_0 + \dots + k_{N-s-1} b_s$	$l_N b_0 + l_{N-1} b_1 + \dots + l_{N-s} b_s$
$s+1$	$k_0 b_0$	$k_1 b_0 + k_0 b_1$	$k_2 b_0 + k_1 b_1 + k_0 b_2$	$\dots k_{j-1} b_0 + k_{j-2} b_1 + \dots + k_{j-s-1} b_s$	$\dots k_{N-2} b_0 + \dots + k_{N-s-2} b_s$	$l_{N-1} b_0 + l_{N-2} b_1 + \dots + l_{N-s-1} b_s$
$s+2$	\vdots	\vdots	\vdots	\dots	\vdots	\vdots
$2s$	\vdots	\vdots	\vdots	$\dots k_{j-s} b_0 + k_{j-s-1} b_1 + \dots + k_{j-2s} b_s$	\dots	$l_{N-s} b_0 + l_{N-s-1} b_1 + \dots + l_{N-2s} b_s$
$N-1$	\vdots	\vdots	\vdots	$k_{j-N+s+1} b_0 + \dots + k_{j-N+1} b_s$	$k_s b_0 + \dots + k_0 b_s$	$l_{s+1} b_0 + l_s b_1 + \dots + l_1 b_s$
N	\vdots	\vdots	\vdots	$k_{j-N+s} b_0 + \dots + k_{j-N} b_s$	$\dots k_{s-1} b_0 + \dots + k_0 b_{s-1}$	$l_s b_0 + l_{s-1} b_1 + \dots + l_0 b_s$

where $l_r = k_r + k_{r+1} + \dots$

$$l_r = k_r + k_{r+1} + k_{r+2} + \dots \tag{5}$$

In short we have

$$p_j = \sum_{i=0}^N r_{ij} p_i \quad j = 0, 1, 2, \dots, N$$

where r_{ij} are the elements in the transition matrix. We shall consider the above expression only for $j \leq N - 1$, since for $j = N$, we have from (5),

$$p_N = \frac{l_N p_0 + (l_N B_{s-1} + l_{N-1} b_s) p_1 + \dots + (l_{s-1} b_0 + l_s b_1 + \dots + l_1 b_s) p_{N-1}}{\left(1 - \sum_{r=0}^s l_r b_{s-r}\right)} \tag{6}$$

Therefore

$$\begin{aligned} p_j &= \sum_{i=0}^N r_{ij} p_i \quad (j = 0, 1, 2, \dots, N-1) \\ &= \sum_{i=0}^{s-1} r_{ij} p_i + \sum_{i=s}^N r_{ij} p_i = A_1 + A_2 \quad (\text{say}) \end{aligned} \tag{7}$$

where

$$\begin{aligned} A_1 &= \sum_{i=0}^{s-1} r_{ij} p_i = k_j p_0 \\ &\quad + (k_j B_{s-1} + k_{j-1} b_s) p_1 \\ &\quad + (k_j B_{s-2} + k_{j-1} b_{s-1} + k_{j-2} b_s) p_2 \\ &\quad + \dots \\ &\quad + (k_j B_1 + k_{j-1} b_2 + \dots + k_{j-s+1} b_s) p_{s-1} \\ &= k_j \sum_{i=0}^{s-1} B_{s-i} p_i + \sum_{i=1}^{s-1} p_i \sum_{r=s-i+1}^s k_{j-i+s-r} b_r \end{aligned} \tag{8}$$

$$\begin{aligned} A_2 &= \sum_{i=s}^N r_{ij} p_i = (k_j b_0 + k_{j-1} b_1 + \dots + k_{j-s} b_s) p_s \\ &\quad + \dots \\ &\quad + (k_{j-N+s} b_0 + \dots + k_{j-N} b_s) p_N \\ &= \sum_{i=s}^N p_i \sum_{r=0}^s k_{j-i+s-r} b_r \end{aligned} \tag{9}$$

Using these values of A_1 and A_2 in (7), we have

$$p_j = k_j \sum_{i=0}^{s-1} B_{s-i} p_i + \sum_{i=1}^{s-1} p_i \sum_{r=s-i+1}^s k_{j-i+s-r} b_r + \sum_{i=s}^N p_i \sum_{r=0}^s k_{j-i+s-r} b_r$$

Multiplying the above expression by x^j and summing over $j=0, 1, 2, \dots, N-1$ and using (1) and (3), we have

$$\begin{aligned} P(x) &= K(x) \sum_{i=0}^{s-1} B_{s-i} p_i + \sum_{j=0}^{N-1} x^j \sum_{i=1}^{s-1} p_i \sum_{r=s-i+1}^s k_{j-i+s-r} b_r \\ &\quad + \sum_{j=0}^{N-1} x^j \sum_{i=s}^N p_i \sum_{r=0}^s k_{j-i+s-r} b_r = C_1 + C_2 + C_3 \quad (\text{say}) \end{aligned} \tag{10}$$

where

$$C_1 = K(x) \sum_{i=0}^{s-1} B_{s-i} p_i$$

$$C_2 = \sum_{j=0}^{N-1} x^j \sum_{i=1}^{s-1} p_i \sum_{r=s-i+1}^s k_{j-i+s-r} b_r$$

Algebraic simplification of this term gives

$$\begin{aligned} C_2 &= (K(x) \cdot x^{-s} \sum_{i=1}^{s-1} \sum_{r=s-i+1}^s p_i b_r x^{i+r} - x^N \sum_{i=1}^{s-r-1} \sum_{r=0}^{s-2} \sum_{j=1}^i p_{i+r} b_{s-r} k_{N-j} x^{i-j}) \\ &= \frac{K(x)}{x^s} \sum_{i=1}^{s-1} p_i x^i \sum_{r=s-i+1}^s b_r x^r - x^N \cdot C^* \end{aligned} \tag{11}$$

where

$$C^* = \sum_{r=0}^{s-2} \sum_{i=1}^{s-r-1} \sum_{j=1}^i p_{i+r} b_{s-r} k_{N-j} x^{i-j} = \sum_{r=0}^{s-2} \sum_{i=r+1}^{s-1} \sum_{j=1}^{i-r} p_i b_{s-r} k_{N-j} x^{i-r-j} \tag{12}$$

$$C_3 = \sum_{j=0}^{N-1} x^j \sum_{i=s}^N p_i \sum_{r=0}^s k_{j-i+s-r} b_r = \sum_{i=s}^N \sum_{j=i-s}^{N-1} \sum_{r=0}^s p_i k_{j-i+s-r} b_r x^j \tag{13}$$

Let $m = j - i + s - r$

Then

$$\begin{aligned} C_3 &= \sum_{i=s}^N \sum_{m=-r}^{N-1-i+s-r} \sum_{r=0}^s p_i k_m b_r x^{m+i-s+r} \\ &= \sum_{i=s}^N p_i x^{i-s} \sum_{m=-r}^{N-1-i+s-r} k_m x^m \sum_{r=0}^s b_r x^r \\ &= \sum_{i=s}^N p_i x^{i-s} \sum_{r=0}^s \sum_{m=-r}^{N-1-i+s-r} k_m x^m \cdot b_r x^r \end{aligned}$$

Thus

$$\begin{aligned} C_3 &= \sum_{i=s}^N p_i x^{i-s} \sum_{r=0}^s b_r \sum_{m=0}^{N-1-i+s-r} x^r k_m x^m, \text{ since } k_j = 0 \text{ for } j < 0. \\ &= \sum_{i=s}^N p_i x^{i-s} \sum_{r=0}^s b_r x^r \left[\sum_{m=0}^{N-1} k_m x^m - \sum_{m=N-i+s-r}^{N-1} k_m x^m \right] \\ &= \sum_{i=s}^N p_i x^{i-s} B_s(x) \cdot K(x) - \sum_{i=s}^N p_i x^{i-s} \sum_{r=0}^s b_r \sum_{m=N-i+s-r}^{N-1} k_m x^{m+r} \\ &= \frac{K(x)}{x^s} \sum_{i=s}^N p_i x^i B_s(x) - \sum_{i=s}^N \sum_{r=0}^s b_r \sum_{m=N-i+s-r}^{N-1} p_i k_m x^{i+m-s+r} \end{aligned} \tag{14}$$

Now substituting the values of C_1 , C_2 and C_3 in (10), we have

$$\begin{aligned} P(x) &= K(x) \sum_{i=0}^{s-1} B_{s-i} p_i + \frac{K(x)}{x^s} \sum_{i=1}^{s-1} p_i x^i \sum_{r=s-i+1}^s b_r x^r - x^N \cdot C^* \\ &\quad + \frac{K(x)}{x^s} \sum_{i=s}^N p_i x^i B_s(x) - \sum_{r=0}^s b_r \sum_{i=s}^N \sum_{m=N-i+s-r}^{N-1} p_i k_m x^{i+m-s+r} \\ &= K(x) \sum_{i=0}^{s-1} B_{s-i} p_i + \frac{K(x)}{x^s} \left[\sum_{i=1}^{s-1} p_i x^i \sum_{r=s-i+1}^s b_r x^r + \sum_{i=s}^N p_i x^i B_s(x) \right] \\ &\quad - x^N \cdot C^* - \sum_{r=0}^s b_r \sum_{i=s}^N \sum_{m=N-i+s-r}^{N-1} p_i k_m x^{i+m-s+r} \end{aligned} \tag{15}$$

Algebraic simplification of the terms in braces on the right hand side in the above expression will give us

$$\begin{aligned}
 P(x) &= K(x) \sum_{i=0}^{s-1} B_{s-i} p_i + \frac{K(x)}{x^s} \left[\sum_{i=0}^N p_i x^i \sum_{r=0}^s b_r x^r - \sum_{i=0}^{s-1} p_i x^i \sum_{r=0}^{s-i} b_r x^r \right] \\
 &\quad - x^N \cdot C^* - \sum_{r=0}^s b_r \sum_{i=s}^N \sum_{m=N-i+s-r}^{N-1} p_i k_m x^{m+i-s+r} \\
 &= K(x) \sum_{i=0}^{s-1} B_{s-i} p_i + \frac{K(x)}{x^s} [\{P(x) + p_N x^N\} B_s(x) \\
 &\quad - \sum_{i=0}^{s-1} p_i x^i B_{s-i}(x)] - x^N \cdot C^* - \sum_{r=0}^s b_r \sum_{i=s}^N \sum_{m=N-i+s-r}^{N-1} p_i k_m x^{m+i-s+r}
 \end{aligned}$$

Thus

$$\begin{aligned}
 P(x) \left[1 - \frac{K(x)}{x^s} B_s(x) \right] &= K(x) \left[\sum_{i=0}^{s-1} B_{s-i} p_i - \sum_{i=0}^{s-1} p_i x^{i-s} B_{s-i}(x) \right] \\
 &\quad - x^N \left[C^* - p_N B_s(x) \cdot \frac{K(x)}{x^s} \right] - \sum_{r=0}^s b_r \sum_{i=s}^N \sum_{m=N-i+s-r}^{N-1} p_i k_m x^{m+i-s+r}
 \end{aligned}$$

Therefore

$$\begin{aligned}
 P(x) &= \frac{K(x) \sum_{i=0}^{s-1} p_i [B_{s-i} x^s - x^i B_{s-i}(x)]}{x^s - K(x) B_s(x)} \\
 &\quad + x^N \left[\frac{C^* x^s - p_N B_s(x) \cdot K(x)}{K(x) B_s(x) - x^s} \right] + \frac{\sum_{i=s}^N \sum_{r=0}^s \sum_{m=N-i+s-r}^{N-1} b_r p_i k_m x^{m+i+r}}{K(x) B_s(x) - x^s}
 \end{aligned}$$

where C^* is given by (12). (16)

Only the first term on the right hand side in (16) will contribute to the coefficients of x^j for $j \leq N-1$. We disregard the second and third terms on the right hand side in (16), since they give the coefficients of x^j for $j \geq N$. These are not needed because we wish to compare the coefficients of x^j for $j < N$ on both sides in (16) in order to evaluate p_j for $j < N$. Let

$$\begin{aligned}
 Q(x) &= \frac{K(x) \sum_{i=0}^{s-1} p_i \{x^s B_{s-i} - x^i B_{s-i}(x)\}}{x^s - K(x) B_s(x)} \\
 &= \frac{\sum_{i=0}^{s-1} p_i \{x^s B_{s-i} - x^i B_{s-i}(x)\}}{x^s / K(x) - B_s(x)}
 \end{aligned}
 \tag{17}$$

The function $Q(x)$ is fully determined once p_0, p_1, \dots, p_{s-1} are specified. This can be done by applying the usual arguments of analyticity of $Q(x)$ and Rouché's Theorem. The expected value of Q can only be obtained by evaluating the probabilities p_0, p_1, \dots, p_n . It should be noted that the usual method of getting $E(Q)$ as $P'(1)$ can not be used since $P(X)$ is not obtained as a closed form expression.

4. Total Time Spent in the System:

Let W be the total time spent by a departing customer, who leaves Q customers behind. Let

$F(t)$ be its distribution function. Then

$$p_j = \Pr\{Q = j\} = \int_0^\infty e^{-\lambda t} \frac{(\lambda t)^j}{j!} dF(t), \quad j < N$$

$$p_N = \Pr\{Q = N\} = \sum_{j=N}^\infty \int_0^\infty e^{-\lambda t} \frac{(\lambda t)^j}{j!} dF(t)$$

From these two expressions, we have

$$\sum_{j=0}^N j p_j = \sum_{j=0}^{N-1} j \int_0^\infty e^{-\lambda t} \frac{(\lambda t)^j}{j!} dF(t) + N \sum_{j=N}^\infty \int_0^\infty e^{-\lambda t} \frac{(\lambda t)^j}{j!} dF(t)$$

$$\therefore E(Q) = \int_0^\infty \sum_{j=0}^{N-1} j e^{-\lambda t} \frac{(\lambda t)^j}{j!} dF(t) + N \int_0^\infty \sum_{j=N}^\infty e^{-\lambda t} \frac{(\lambda t)^j}{j!} dF(t)$$

then

$$E(Q) = \int_0^\infty \sum_{j=0}^{N-1} j e^{-\lambda t} \frac{(\lambda t)^j}{j!} dF(t) + N \int_0^\infty \left[1 - \sum_{j=0}^{N-1} e^{-\lambda t} \frac{(\lambda t)^j}{j!} \right] dF(t)$$

$$\begin{aligned} E(Q) &= N + \int_0^\infty \sum_{j=0}^{N-1} (j-N) e^{-\lambda t} \frac{(\lambda t)^j}{j!} dF(t) \\ &= N + \int_0^\infty \left[\lambda t \sum_{j=0}^{N-2} e^{-\lambda t} \frac{(\lambda t)^j}{j!} - N \sum_{j=0}^{N-1} e^{-\lambda t} \frac{(\lambda t)^j}{j!} \right] dF(t) \\ &= N + \int_0^\infty \left[\lambda t \frac{\Gamma(N-1, \lambda t)}{\Gamma(N-1)} - N \frac{\Gamma(N, \lambda t)}{\Gamma(N)} \right] dF(t) \end{aligned}$$

where

$$\sum_{j=0}^N \frac{e^{-\lambda} \lambda^j}{j!} = \frac{\Gamma(N+1, \lambda)}{\Gamma(N+1)}$$

Thus

$$E(Q) = N + \frac{1}{\Gamma(N)} \int_0^\infty [(N-1)\lambda t \Gamma(N-1, \lambda t) - N\Gamma(N, \lambda t)] dF(t)$$

or

$$\int_0^\infty [N\Gamma(N, \lambda t) - (N-1)\lambda t \Gamma(N-1, \lambda t)] dF(t) = \Gamma(N)[N - E(Q)]$$

This expression relates the waiting time distribution and the average number of customers in the system. Thus the waiting time can be found by using the properties of incomplete gamma functions from the above expressions.

5. Special Cases*

Acknowledgements

The author wishes to express his sincere appreciation to Dr. U. N. Bhat for his helpful comments and suggestions.

* This section will be published as an Addendum to this paper in the next issue of this journal, Volume 6, nr. 1, 1972.

REFERENCES

- [1] U. N. Bhat, Customer Overflow in Queues with Finite Waiting Space, *Australian Journal of Statistics*, 7, No. 1, (1965) 15–19.
- [2] U. N. Bhat, Imbedded Markov Chains in Queuing System M/G/1 and GI/M/1 with Limited Waiting Room, *Research Report, Department of Statistics*, Michigan State University, Feb 25, 1966. (Abstract appears in *AMS*, 37, No. 2 (1966) 540.
- [3] P. D. Finch, The Effect of the Size of the Waiting Room on a Simple Queue, *Journal of the Royal Statistical Society, Series B*, 20, No. 1 (1958) 182–186.
- [4] H. C. Jain, Queuing Problem with Limited Waiting Space, *Naval Research Logistics Quarterly*, 9, No. 3 (1962) 245–252.
- [5] N. K. Jaiswal, On Some Waiting Line Problems, *Opsearch*, 2, No. 1 and 2 (1965) 27–43.
- [6] J. Keilson, The Ergodic Queue Length Distribution for Queuing Systems with Finite Capacity, *Journal of the Royal Statistical Society, Series C*, Parts 2 and 3 (1965) 190–201.
- [7] S. S. Rao, Queuing Models with Balking, Reneging and Interruptions, *Operations Research*, 13, No. 4 (1965) 596–608.
- [8] V. P. Singh, Finite Storage Problems via Queuing Theory, *Master's Thesis*, Ohio State University, 1967.
- [9] S. Takamatsu, On the Come-and-Stay Interarrival Time in a Modified Queuing System M/G/1, *Ann. Inst. Statist. Math.*, 15 (1963) 73–78.

Journal of Engineering Math., Vol. 5 (1971) 241–248